

# Fixed points of pattern-avoiding involutions

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Joint work with  
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and  
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Permutation Patterns, June 2017, Reykjavik University

## Outline of talk:

- ▶ Analytic combinatorics with bivariate generating functions.
- ▶ Standard Young tableaux and involutions avoiding monotone patterns.

Notation:

- ▶  $\mathbf{Av}(\rho)$  ( $\mathbf{Av}_n(\rho)$ ) =  $\rho$ -avoiding permutations (of size  $n$ ).
- ▶  $\mathbf{Iv}(\rho)$  ( $\mathbf{Iv}_n(\rho)$ ) =  $\rho$ -avoiding involutions (of size  $n$ ).
- ▶  $fp(\sigma)$  = number of fixed points of  $\sigma$ .
- ▶ For a permutation class  $\mathcal{C}$ , the bivariate generating function wrt fixed points is given by:

$$F_{\mathcal{C}}(x, t) = \sum_{\sigma \in \mathcal{C}} x^{fp(\sigma)} t^{|\sigma|}.$$

From Elizalde (2004), for  $\tau = 321, 132$ , or  $213$ ,

$$F_{\mathbf{Av}(\tau)}(x, t) = \sum_{\sigma \in \mathbf{Av}(\tau)} x^{fp(\sigma)} t^{|\sigma|} = \frac{2}{1 + 2t(1 - x) + \sqrt{1 - 4t}}.$$

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By Analytic Combinatorics (Flajolet and Sedgewick 2009):

$$[t^n]F_{\mathbf{Av}(321)}(x, t) \sim -[t^n]\frac{8}{(3 - x)^2}\sqrt{1 - 4t} \sim \frac{4^{n+1}}{(3 - x)^2\sqrt{\pi n^3}}$$

so

$$\frac{[t^n]F_{\mathbf{Av}(321)}(x, t)}{[t^n]F_{\mathbf{Av}(321)}(1, t)} \rightarrow \frac{4}{(3 - x)^2} = \sum_{k=0}^{\infty} \frac{4}{9}(k + 1) \left(\frac{1}{3}\right)^k x^k.$$

Enumeration involutions avoiding patterns of length 3 (Simion and Schmidt, 1985)

- ▶ For  $\tau = 123, 321, 132, 213$ ,  $|\mathbf{Iv}_n(\tau)| = \binom{n}{\lfloor n/2 \rfloor}$ .
- ▶ For  $\tau = 231, 312$ ,  $|\mathbf{Iv}_n(\tau)| = 2^{n-1}$

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Asymptotic enumeration of involutions avoiding monotone patterns (Regev 1981)

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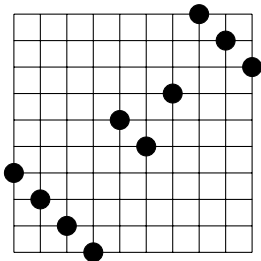
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Involutions avoiding patterns of length 4 (Bóna, Homberger, Pantone, and Vatter 2014)

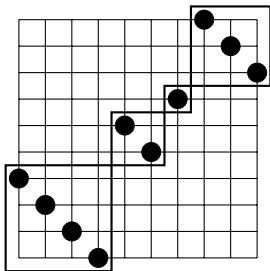


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$$\alpha = (4, 2, 1, 3).$$



$$F_{\mathbf{lv}(231)} = F_{\mathbf{lv}(312)} = \frac{1 - t^2}{1 - 2t^2 - xt}.$$

## Theorem (Miner, Rizzolo, S. 2017)

Let  $\Pi_n$  denote a uniform random element from  $\mathbf{lv}_n(231) = \mathbf{lv}_n(312)$ . Then

$$\frac{fp(\Pi_n) - \frac{1}{3}n}{\sqrt{8n/27}} \rightarrow_d Z,$$

where  $Z$  is a standard normal random variable.

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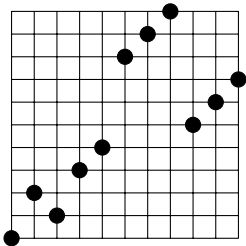
Prove using Theorem IX.9 of Flajolet and Sedgewick on

$$F_{\mathbf{lv}(231)} = \frac{1 - t^2}{1 - 2t^2 - xt}.$$

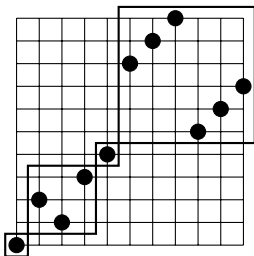
$$\mathbb{E}(fp(\Pi_n)) = (n/3) + O(1).$$

$$\text{Var}(fp(\Pi_n)) = (8n/27) + O(1).$$

$$\sigma = 1\ 3\ 2\ 4\ 5\ 9\ 10\ 11\ 6\ 7\ 8 \in \mathbf{lv}_{11}(321).$$

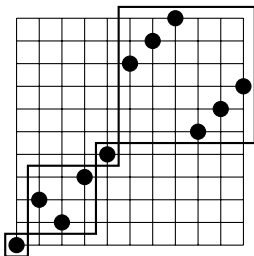


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If  $\Pi_n$  is a uniformly random element of  $\mathbf{lv}_n(\tau)$ , then

$$\sqrt{\frac{1}{n}}fp(\Pi_n) \rightarrow_d X,$$

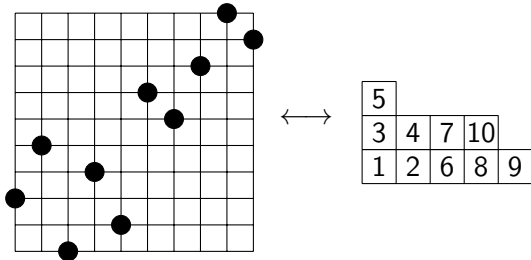
where  $X$  follows a Rayleigh(1) distribution with density function  $f(x) = xe^{-x^2/2}$ .



## Avoiding monotone patterns and standard Young Tableaux

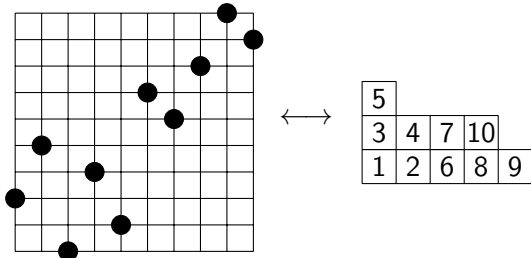
## Classic Results

- ▶  $\tau_k = (k+1)k \cdots 21$ ,  $\rho_k = 12 \cdots k(k+1)$ .



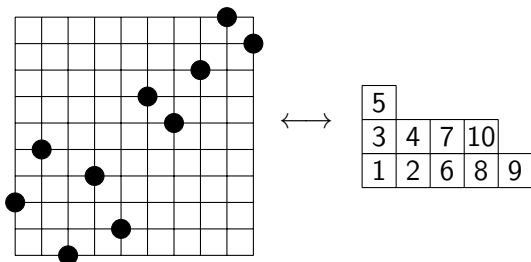
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- ▶  $\mathbf{lv}_n(\tau_k)$  (or  $\mathbf{lv}_n(\rho_k)$ ) is in bijection with standard Young tableaux with at most  $k$  rows (or  $k$  columns) through RSK.  
[Schensted, 1961]



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[Schensted, 1961]
- ▶ Fixed points of  $\pi$  = number of odd columns.  
[Schutzenberger, 1963]



## Theorem (Matsumoto 2008)

Let  $(\Lambda_i)_{1 \leq i \leq k}$  be the ranked eigenvalues of a traceless  $k \times k$  GOE matrix. For an involution  $\pi$ , let  $\lambda_i(\pi)$  denote the length of the  $i$ th row of the tableaux of  $\pi$  under RSK. For  $\Pi_n$  chosen uniformly from  $\mathbf{Iv}_n(\tau_k)$ ,

$$\left( \sqrt{\frac{k}{n}} \left( \lambda_i(\Pi_n) - \frac{n}{k} \right) \right)_{1 \leq i \leq k} \longrightarrow_d (\Lambda_i)_{1 \leq i \leq k}.$$

Moreover, for any fixed  $d > 0$ ,

$$\mathbb{P} \left( \min_{2 \leq i \leq k} \{ \lambda_{i-1}(\Pi_n) - \lambda_i(\Pi_n) \} < d \right) \rightarrow 0.$$

## Theorem (Miner, Rizzolo, S. 2017)

Fix  $k \in \{2, 3, \dots\}$ . Let  $(\Lambda_i)_{1 \leq i \leq k}$  be the ranked eigenvalues of a traceless  $k \times k$  GOE matrix and let  $\Pi_n$  be a uniformly random element of  $\mathbf{Iv}_n((k+1)k \cdots 321)$ .

1. If  $k$  is even then

$$\sqrt{\frac{k}{n}} f p(\Pi_n) \rightarrow_d \sum_{j=1}^k (-1)^{j+1} \Lambda_j.$$

2. If  $k$  is odd then

$$\sqrt{\frac{k}{n}} \left( f p(\Pi_n) - \frac{n}{k} \right) \rightarrow_d \sum_{j=1}^k (-1)^{j+1} \Lambda_j.$$

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Note  $fp(\Pi_n) = \sum_{j=1} \lambda_{2j-1}(\Pi_n) - \lambda_{2j}(\Pi_n)$ .

## Theorem (Miner, Rizzolo, S. 2017)

If  $\Pi_n$  is a uniformly random element in  $\mathbf{lv}_n(123 \cdots k(k+1))$  then

$$fp(\Pi_{2n}) \rightarrow_d X_{\text{even}}$$

and

$$fp(\Pi_{2n-1}) \rightarrow_d X_{\text{odd}},$$

where  $X_{\text{even}}$  has density function given by

$$\mathbb{P}(X_{\text{even}} = i) = \begin{cases} \frac{\binom{k}{i}}{2^{k-1}} & i \text{ is even}, \\ 0 & i \text{ is odd}, \end{cases}$$

and  $X_{\text{odd}}$  has density function given by

$$\mathbb{P}(X_{\text{odd}} = i) = \begin{cases} \frac{\binom{k}{i}}{2^{k-1}} & i \text{ is odd}, \\ 0 & i \text{ is even}. \end{cases}$$



Markov chain  $C$  with state space  $S = \{0, 1, \dots, k\}$ , and transition matrix  $P$  with probabilities

$$P_{i,j} = \begin{cases} \frac{i}{k} & j = i - 1 \\ 1 - \frac{i}{k} & j = i + 1 \\ 0 & \text{otherwise.} \end{cases}$$

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### Lemma

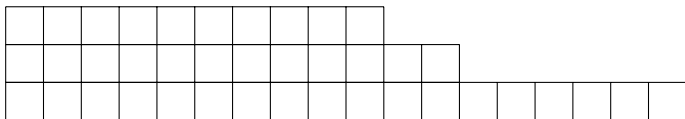
*The chain  $C$  is periodic of period 2. As  $d \rightarrow \infty$ ,  $C_d$  approaches alternation between vectors  $\mathbf{p}$  and  $\mathbf{q} \in S$ , where*

$$\mathbf{p}_i = \begin{cases} \frac{\binom{k}{i}}{2^{k-1}} & i \text{ is even} \\ 0 & \text{otherwise,} \end{cases}$$

*and*

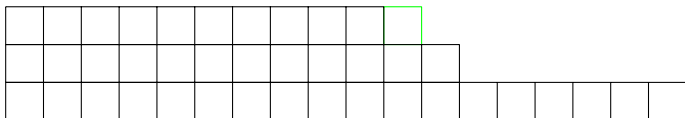
$$\mathbf{q}_i = \begin{cases} \frac{\binom{k}{i}}{2^{k-1}} & i \text{ is odd} \\ 0 & \text{otherwise.} \end{cases}$$

Fix  $d > 0$  and pick a Young diagram of size  $n - d$  with  $k$  rows



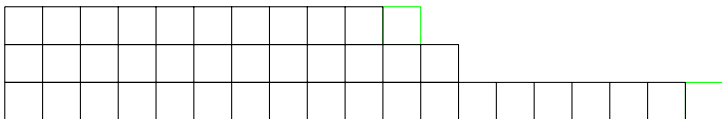
If minimum separation is less than  $d$  total number of Young diagrams of size  $n$  containing the smaller is  $k^d$ .

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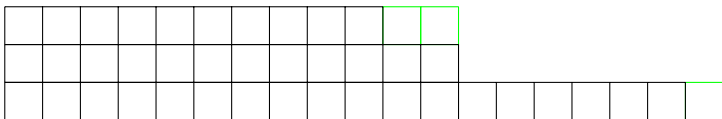
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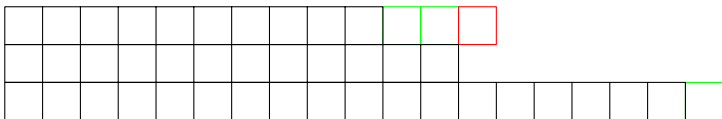
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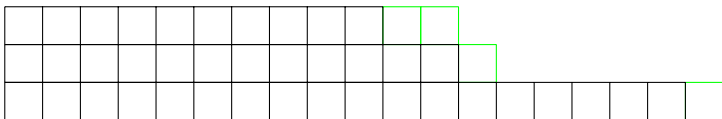
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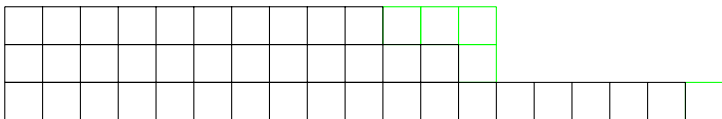
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Weights on Young diagrams  $\lambda$  of size  $n$ .

- ▶  $f_\lambda$  = number of standard fillings of  $\lambda$
- ▶  $s_\lambda^d$  = number of ways to fill  $d$  extra boxes after filling  $\lambda$ .

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Two measures on standard Young diagrams:

$$\mu_1 = \frac{f_\lambda}{\sum_\gamma f_\gamma} \quad \text{and} \quad \mu_2 = \frac{f_\lambda s_\lambda^d}{\sum_\gamma f_\gamma s_\gamma^d}.$$

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By Matsumoto  $\mathbb{P}(s_\lambda^d < k^d) \rightarrow 0$ .

Thanks!